

Bayesian Tools for Econometric Analysis

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Part 3: Bayesian treatment of the univariate linear regression model (ULRM)

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1 The conjugate prior analysis

1.1 The known variance case

Suppose we have T observations on the k covariate regression model with NID Gaussian errors

$$\underset{(T \times 1)}{\mathbf{y}} = \underset{(T \times k)}{\mathbf{X}} \underset{(k \times 1)}{\boldsymbol{\beta}} + \underset{(T \times 1)}{\boldsymbol{\varepsilon}} \quad (1)$$

$$\boldsymbol{\varepsilon} \sim N(\mathbf{0}, \sigma^2 \mathbf{I}_T) \quad (2)$$

with $h = \sigma^{-2}$, error precision, known and \mathbf{X} exogenous.

Let us focus on the Bayesian estimation of $\boldsymbol{\beta}$.

Non-Bayesians:

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} \quad (3)$$

$$V(\hat{\beta}) = \sigma^2(\mathbf{X}'\mathbf{X})^{-1} \quad (4)$$

Likelihood:

$$p(\mathbf{y}|\mathbf{X}, \beta, \sigma^2) = (\sqrt{2\pi})^{-T} (\sigma^2)^{-T/2} \exp\left\{-\frac{1}{2\sigma^2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon}\right\} \quad (5)$$

$$\boldsymbol{\varepsilon} = \mathbf{y} - \mathbf{X}\beta \quad (6)$$

Prior for β :

$$p(\beta) = N \left(\begin{array}{c} \underline{\beta} \\ (k \times 1) \end{array}, \begin{array}{c} \underline{\mathbf{H}}^{-1} \\ (k \times k) \end{array} \right) = (2\pi)^{-k/2} |\underline{\mathbf{H}}|^{1/2} \times \\ \times \exp \left\{ -\frac{1}{2} (\beta - \underline{\beta})' \underline{\mathbf{H}}_{\beta} (\beta - \underline{\beta}) \right\}$$

- absence of a priori info: $\underline{\mathbf{H}} = [\mathbf{0}]$

- dogmatic prior: $\underline{\mathbf{H}}^{-1} = [\mathbf{0}]$

⇒ Bayes Theorem:

$$\begin{aligned} p(\beta|\mathbf{y}, \mathbf{X}, \sigma^2) &\propto \exp \left\{ -\frac{1}{2} (\beta - \underline{\beta})' \underline{\mathbf{H}}_{\beta} (\beta - \underline{\beta}) \right\} \\ &\times \exp \left\{ -\frac{1}{2\sigma^2} \varepsilon' \varepsilon \right\} \\ &\propto \exp \left\{ -\frac{1}{2} \left[(\beta - \underline{\beta})' \underline{\mathbf{H}}_{\beta} (\beta - \underline{\beta}) + \frac{1}{\sigma^2} \varepsilon' \varepsilon \right] \right\} \end{aligned} \quad (7)$$

Remember OLS decomposition

$$\varepsilon' \varepsilon = (\beta - \hat{\beta})' \mathbf{X}' \mathbf{X} (\beta - \hat{\beta}) + \hat{\varepsilon}' \hat{\varepsilon} \quad (8)$$

$$\begin{aligned} \implies \exp\left(-\frac{1}{2\sigma^2} \varepsilon' \varepsilon\right) &= \exp\left(-\frac{1}{2\sigma^2} \hat{\varepsilon}' \hat{\varepsilon}\right) \times \\ &\exp\left(-\frac{1}{2\sigma^2} (\beta - \hat{\beta})' \mathbf{X}' \mathbf{X} (\beta - \hat{\beta})\right) \end{aligned} \quad (9)$$

Clearly Gaussian for β , with moments which can be obtained by writing exponential part as a sum of squares:

$$\begin{aligned} \text{sample info: } \underset{(T \times 1)}{\mathbf{y}} &= \mathbf{X}\beta + \varepsilon, \varepsilon \sim N(\mathbf{0}, \sigma^2 \mathbf{I}_T) \\ \text{a priori info: } \underset{(k \times 1)}{\underline{\beta}} &= \beta + \varepsilon_0, \varepsilon_0 \sim N(\mathbf{0}, \underline{\mathbf{H}}^{-1}) \end{aligned} \quad (10)$$

Standardising:

$$\text{sample: } \frac{\mathbf{1}}{\sigma} \mathbf{y}_{(T \times 1)} = \frac{\mathbf{1}}{\sigma} \mathbf{X} \boldsymbol{\beta} + \mathbf{e}, \mathbf{e} = \frac{\mathbf{1}}{\sigma} \boldsymbol{\varepsilon} \sim \mathbf{N}(\mathbf{0}, \mathbf{I}_T)$$

$$\text{a priori: } \mathbf{P}^{-1} \underline{\boldsymbol{\beta}}_{(k \times 1)} = \mathbf{P}^{-1} \boldsymbol{\beta} + \mathbf{e}_0, \mathbf{e}_0 = \mathbf{P}^{-1} \boldsymbol{\varepsilon}_0, \quad (11)$$

$$\mathbf{e}_0 \sim N(\mathbf{0}, \mathbf{I}_k) \quad (12)$$

$$\underline{\mathbf{H}}^{-1} = \mathbf{P} \mathbf{P}' \quad (13)$$

\Rightarrow

$$\mathbf{y}^*_{(T+k) \times 1} = \begin{bmatrix} \frac{1}{\sigma} \mathbf{y}_{(T \times 1)} \\ \mathbf{P}^{-1} \underline{\mu}_\beta \\ \mathbf{P}^{-1} \mathbf{e}_0 \end{bmatrix}, \quad \mathbf{X}^*_{(T+k) \times k} = \begin{bmatrix} \frac{1}{\sigma} \mathbf{X} \\ \mathbf{P}^{-1} \end{bmatrix}, \quad \mathbf{e}^* = \begin{bmatrix} \mathbf{e} \\ \mathbf{e}_0 \end{bmatrix} \quad (14)$$

$$\mathbf{y}^* = \mathbf{X}^* \beta + \mathbf{e}^*, \quad \mathbf{e}^* \sim N(\mathbf{0}, \mathbf{I}_{T+k}) \quad (15)$$

$$\begin{aligned} \bar{\beta} &= (\mathbf{X}^{*'} \mathbf{X}^*)^{-1} \mathbf{X}^{*'} \mathbf{y}^* = \left(\frac{1}{\sigma^2} \mathbf{X}' \mathbf{X} + \mathbf{P}^{-1'} \mathbf{P}^{-1} \right)^{-1} \\ &\quad \times \left(\frac{1}{\sigma^2} \mathbf{X}' \mathbf{y} + \mathbf{P}^{-1'} \mathbf{P}^{-1} \underline{\beta} \right) \\ &= \left(\frac{1}{\sigma^2} \mathbf{X}' \mathbf{X} + \underline{\mathbf{H}} \right)^{-1} \left(\frac{1}{\sigma^2} \mathbf{X}' \mathbf{y} + \underline{\mathbf{H}} \underline{\beta} \right) = \\ &= (h \mathbf{X}' \mathbf{X} + \underline{\mathbf{H}})^{-1} (h \mathbf{X}' \mathbf{y} + \underline{\mathbf{H}} \underline{\beta}) \end{aligned} \quad (16)$$

$$\bar{\mathbf{H}} = \left(\frac{1}{\sigma^2} \mathbf{X}' \mathbf{X} + \underline{\mathbf{H}} \right) = (h \mathbf{X}' \mathbf{X} + \underline{\mathbf{H}}) \quad (17)$$

Note: inputs for moments of the posterior distribution of β :

1. $\underline{\mathbf{H}}, \underline{\mathbf{H}} \times \underline{\beta}$

2. $\mathbf{X}'\mathbf{X}, \mathbf{X}'\mathbf{y}$

3. h

\Rightarrow output:

1. $\bar{\beta}$

2. $\bar{\mathbf{H}}$

1.2 Unknown variance case

1.2.1 Digression: Gamma distribution

If $s \cdot x$ is Chisquare(ν)

$$p(s \cdot x) = \frac{1}{2^{\frac{\nu}{2}} \Gamma(\nu/2)} (s \cdot x)^{\frac{\nu}{2}-1} \exp(-s \cdot x/2)$$
$$p(x) = \frac{\left(\frac{s}{2}\right)^{\frac{\nu}{2}}}{\Gamma(\nu/2)} (x)^{\frac{\nu}{2}-1} \exp(-s \cdot x/2)$$

$$E(x) = \frac{\nu}{s}$$
$$V(x) = \frac{2\nu}{s^2}$$
$$Mode(x) = \frac{\nu - 2}{s}$$

we say x is Gamma distributed.

In the ULRM:

- MRL: $\mathbf{y} = \mathbf{X}\beta + \varepsilon$
- \mathbf{X} ($T \times k$) exogenous regressors matrix
- $\varepsilon \sim N(\mathbf{0}, h^{-1}\mathbf{I}_T)$
- $h = \frac{1}{\sigma^2}$ precision

Likelihood:

$$p(\mathbf{y}|\mathbf{X}, \beta, \sigma^2) = \left(\sqrt{2\pi}\right)^{-T} h^{T/2} \exp\left\{-\frac{h}{2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon}\right\}$$
$$\boldsymbol{\varepsilon} = \mathbf{y} - \mathbf{X}\beta$$

Conjugate prior:

$$p(\theta) = p(\beta, h) = p(\beta|h) \times p(h)$$
$$p(\beta|h) = N(\underline{\beta}, h^{-1}\underline{\mathbf{H}}),$$
$$p(h) : \underline{s} \times h \sim \chi_{\underline{\nu}}^2 \text{ (} h \text{ Gamma distributed)}$$

posterior distribution

$$\begin{aligned} p(\theta|\mathbf{X}, \mathbf{y}) &\propto p(\theta) \times p(\mathbf{y}|\mathbf{X}, \theta) \\ &= h^{T/2} \exp\left\{-\frac{h}{2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon}\right\} \times \\ &h^{k/2} \exp\left\{-\frac{h}{2}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})'\underline{\mathbf{H}}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})\right\} \times \\ &h^{\frac{\nu-2}{2}} \exp\left(-\frac{1}{2}\underline{s}h\right) \\ &= h^{\frac{T+k+\nu-2}{2}} \exp\left\{-\frac{h}{2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon} - \frac{h}{2}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})'\underline{\mathbf{H}}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}}) - \frac{1}{2}\underline{s}h\right\} \end{aligned}$$

Easy to describe conditional posterior distributions

$$p(\beta|h, \mathbf{X}, \mathbf{y}) = N(\bar{\beta}, \bar{\mathbf{H}}^{-1}) \quad (18)$$

$$p(h|\beta, \mathbf{X}, \mathbf{y}) : \bar{s} \times h \sim \chi_{\bar{\nu}}^2 \quad (19)$$

(similar to known variance case)

$$\bar{\beta} = (\mathbf{X}'\mathbf{X} + \underline{\mathbf{H}})^{-1} (\mathbf{X}'\mathbf{y} + \underline{\mathbf{H}}\underline{\beta}) \quad (20)$$

$$\bar{\mathbf{H}}_{\beta}^{-1} = h^{-1} (\mathbf{X}'\mathbf{X} + \underline{\mathbf{H}})^{-1} \quad (21)$$

$$\bar{s} = [\underline{s} + \varepsilon'\varepsilon] \quad (22)$$

$$\bar{\nu} = \underline{\nu} + T + k \quad (23)$$

and the marginal posterior distribution $p(\beta|\mathbf{X}, \mathbf{y})$?

$$p(\beta|\mathbf{X}, \mathbf{y}) = \int p(\beta, h|\mathbf{X}, \mathbf{y})dh \quad (24)$$

is Student t

Moreover, marginal likelihood is known analytically

$$\begin{aligned}
 p(\mathbf{y}|\mathbf{X}) &= \int p(\beta, h)p(\mathbf{y}|\mathbf{X}, \beta, h)dhd\beta = \\
 &= \pi^{-T/2}\Gamma\left(\frac{T+\nu}{2}\right)\Gamma\left(\frac{\nu}{2}\right)\left[\frac{|\mathbf{H}|}{|\bar{\mathbf{H}}|}\right]^{1/2}\underline{\mathbf{s}}^{\frac{\nu}{2}}\times \\
 &\left[\underline{\mathbf{s}} + \hat{\boldsymbol{\varepsilon}}'\hat{\boldsymbol{\varepsilon}} + (\hat{\boldsymbol{\beta}} - \bar{\boldsymbol{\beta}})'\mathbf{X}'\mathbf{X}(\hat{\boldsymbol{\beta}} - \bar{\boldsymbol{\beta}}) + (\underline{\boldsymbol{\beta}} - \bar{\boldsymbol{\beta}})'\mathbf{H}(\underline{\boldsymbol{\beta}} - \bar{\boldsymbol{\beta}})\right]^{-\frac{T+\nu}{2}} \quad (25)
 \end{aligned}$$

complicated but known analytically (see Geweke, 2005).

2 Other prior distributions

Independent priors

$$\begin{aligned} p(\theta) &= p(\beta, h) = p(\beta) \times p(h) \text{ (ind a priori)} \\ p(\beta|h) &= N(\underline{\beta}, \underline{\mathbf{H}}^{-1}), \\ p(h) : \underline{s} \times h &\sim \chi_{\underline{\nu}}^2 \text{ (} h \text{ Gamma distributed)} \end{aligned}$$

Posterior distribution is

$$\begin{aligned} p(\theta|\mathbf{X}, \mathbf{y}) &\propto p(\theta) \times p(\mathbf{y}|\mathbf{X}, \theta) \\ &= h^{T/2} \exp\left\{-\frac{h}{2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon}\right\} \times \\ &\exp\left\{-\frac{1}{2}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})'\underline{\mathbf{H}}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})\right\} \times \\ &h^{\frac{\nu-2}{2}} \exp\left(-\frac{1}{2}\underline{s}h\right) \\ &= h^{\frac{T+\nu-2}{2}} \exp\left\{-\frac{h}{2}\boldsymbol{\varepsilon}'\boldsymbol{\varepsilon} - \frac{1}{2}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}})'\underline{\mathbf{H}}(\boldsymbol{\beta}-\underline{\boldsymbol{\beta}}) - \frac{1}{2}\underline{s}h\right\} \end{aligned}$$

Full set of conditionals

$$p(\beta|h, \mathbf{X}, \mathbf{y}) = N(\bar{\beta}, \bar{\mathbf{H}}^{-1}) \quad (26)$$

$$p(h|\beta, \mathbf{X}, \mathbf{y}) : \bar{s} \times h \sim \chi_{\bar{\nu}}^2 \quad (27)$$

with

$$\bar{\beta} = (h\mathbf{X}'\mathbf{X} + \underline{\mathbf{H}})^{-1} (h\mathbf{X}'\mathbf{y} + \underline{\mathbf{H}}\underline{\beta}) \quad (28)$$

$$\bar{\mathbf{H}}^{-1} = (h\mathbf{X}'\mathbf{X} + \underline{\mathbf{H}})^{-1} \quad (29)$$

$$\bar{s} = [\underline{s} + \varepsilon'\varepsilon] \quad (30)$$

$$\bar{\nu} = \underline{\nu} + T \quad (31)$$

What is marginal distribution of $p(\beta|\mathbf{X}, \mathbf{y})$?

$$p(\beta|\mathbf{X}, \mathbf{y}) = \int p(\beta, h|\mathbf{X}, \mathbf{y})dh \quad (32)$$

not known analytically.

Need to use simulation based methods

Also to obtain marginal likelihood

3 MCMC in the ULRM

Simplest approach: use a Gibbs sampling algorithm, with two blocks:

$$\theta = \begin{bmatrix} \beta \\ h \end{bmatrix}$$

Sequentially sample from:

$$p(\beta^{(i)} | h^{(i-1)}, \mathbf{y}, \mathbf{X}) \tag{33}$$

$$p(h^{(i)} | \beta^{(i)}, \mathbf{y}, \mathbf{X}) \tag{34}$$

from arbitrary initialisation.

This is a Markov chain that converges to posterior distribution .

3.1 Example: a Phillips curve for the US

Amisano and Giacomini (2007)

$$\pi_t = \beta_0 + \sum_{i=1}^4 \beta_i \pi_{t-i} + \gamma_1 u_{t-1} + h^{-1/2} e_t \quad (35)$$

see codes for conjugate and non-conjugate analysis.

The codes are called

`main_linear_regression_PC_01.m` and

`main_linear_regression_PC_02.m`,

respectively.

4 Priors as stochastic constraints

Priors as stochastic constraints

$$\mathbf{R}_{(q \times k)} \beta = \delta + \varepsilon_0 \quad (36)$$

$$\varepsilon_0 \sim N(\mathbf{0}, \underline{\mathbf{H}}^{-1}) \quad (37)$$

if $\underline{\mathbf{H}}$ goes to zero, priors become less and less informative.

Write again in dummy observation form

$$\mathbf{y}^* = \mathbf{X}^* \boldsymbol{\beta} + \mathbf{e}^* \quad (38)$$

$$\mathbf{y}^* = \begin{bmatrix} h^{1/2} \mathbf{y} \\ \underline{\mathbf{H}}^{1/2} \boldsymbol{\delta} \end{bmatrix}, \mathbf{X}^* = \begin{bmatrix} h^{1/2} \mathbf{X} \\ \underline{\mathbf{H}}^{1/2} \mathbf{R} \end{bmatrix}, \mathbf{e}^* = \begin{bmatrix} h^{1/2} \boldsymbol{\varepsilon} \\ \underline{\mathbf{H}}^{1/2} \boldsymbol{\varepsilon}_0 \end{bmatrix} \quad (39)$$

with Gamma prior on h .

→

$$p(\beta|h, \mathbf{y}, \mathbf{X}) = N(\bar{\beta}, \bar{\mathbf{H}})$$
$$\bar{\mathbf{H}} = [h\mathbf{X}'\mathbf{X} + \mathbf{R}'\underline{\mathbf{H}}\mathbf{R}]^{-1}$$
$$\bar{\beta} = \bar{\mathbf{H}} [h\mathbf{X}'\mathbf{y} + \mathbf{R}'\underline{\mathbf{H}}\delta]$$

Theil and Goldberger (1964) mixed estimator.

Examples of this dummy prior

- sum of coefficient priors : VARs, integration and cointegration
- shrinkage priors (Bayesian VARs);
- Del Negro and Schorfheide (IER, 2004): DSGE based prior

4.1 Sum of coefficients priors

Example with sum of coefficient priors: vertical Phillips curve

$$\sum_{i=1}^4 \beta_i = 1 + \textit{noise} \quad (40)$$

4.2 Shrinkage priors

When large parameter space, it is convenient to impose priors that shrink parameter space in a stochastic manner

$$\begin{aligned}\sqrt{V(a_{ij,k})} &= \lambda_0 \times k^{\lambda_1}, \lambda_1 < 0; \\ \lambda_0 &= \textit{overall shrinkage} \\ \lambda_1 &= \textit{lag decay} \\ \sqrt{V(\textit{intercept})} &= \lambda_2\end{aligned}$$

Example with "Taylor Rule"

$$r_t = \beta_0 + \sum_{i=1}^{12} (\beta_{1i}r_{t-i} + \beta_{2i}\pi_{t-i} + \beta_{3i}u_{t-i}) + h^{-1/2}e_t \quad (41)$$

see code `main_linear_regression_TR_01.m`

5 Forecasting

Statement about the future (out of sample). What is unknown about the future?

- validity of the model
- parameters
- future shocks

In a frequentist approach

$$p(y_{T+1}|\mathbf{D}_T, \hat{\theta}_T, M_j) \quad (42)$$

In a Bayesian framework

$$p(y_{T+1}|\mathbf{D}_T, M_j) = \int p(y_{T+1}|\mathbf{D}_T, \theta, M_j)p(\theta|\mathbf{D}_T, M_j)d\theta \quad (43)$$

(marginalise parameters out)

$$p(y_{T+1}|\mathbf{D}_T) = \sum_{j=1}^J p(y_{T+1}|\mathbf{D}_T, M_j)p(M_j|\mathbf{D}_T) \quad (44)$$

(marginalise model out)

$$p(M_j|\mathbf{D}_T) = \frac{p(\mathbf{D}_T|M_j)p(M_j)}{\sum_{i=1}^J p(\mathbf{D}_T|M_i)p(M_i)}$$

Classical forecasts

1. Point forecasts
2. interval forecasts
3. density forecasts